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Assessment of the ORH 3B orange roughy fishery for the 1994-95 fishing year

R. I. C. C. Francis, M. R. Clark, R. P. Coburn, K. D. Field, and P. J. Grimes MAF Fisheries Greta Point<br>P.O. Box 297<br>Wellington

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This series documents the scientific basis for stock assessments and fisheries management advice in New Zealand. It addresses the issues of the day in the current legislative context and in the time frames required. The documents it contains are not intended as definitive statements on the subjects addressed but rather as progress reports on ongoing investigations.

# ASSESSMENT OF THE ORH 3B ORANGE ROUGHY FISHERY FOR THE 1994-95 FISHING YEAR 

R.I.C.C. Francis, M.R. Clark, R.P.Coburn, K.D. Field, and P.J. Grimes<br>N.Z. Fisheries Assessment Research Document 95/4. 43 p.

## 1 Executive Summary

Orange roughy in ORH 3B are assumed to come from four stocks. These are associated with the Chatham Rise, the Waitaki area, the area just south of Puysegur, and the Macquarie Ridge.

### 1.1 Chatham Rise

The current status and virgin biomass of the Chatham Rise stock are estimated using stock reduction analysis. The main inputs to this analysis are the history of catches from the Chatham Rise, and estimates of biomass and mean length from a series of eight trawl surveys in the main spawning area.

It is possible that there is more than one stock on the Chatham Rise because there are significant differences between genetic samples taken from different locations. However, similar differences have been found between samples taken from the same location at different times, so there are no clear stock boundaries within the Rise. There is no reason to believe that the existence of more than one stock in this area would imply either a more optimistic or more pessimistic picture than the current assessment.

Two changes in life history parameters (a change in the maturity ogive and an increase in natural mortality) result in substantial increases in the estimated productivity of orange roughy. However, there remains considerable uncertainty about the true values of these parameters because of uncertainty about age estimates.

The biomass index from the 1994 spawning ground (Box) survey is highly uncertain because $95 \%$ of the biomass is associated with one stratum, and $66 \%$ is associated with a single catch, and there is a strong imbalance in the sex ratio ( $86 \%$ of the estimated biomass is female). Consequently, the analysis was carried out with four alternative assumptions concerning the survey data.

For all four assumptions the stock is estimated to be below the level that will support the MSY. The estimated mid-season biomass for 1993-94 is $10-17 \% B_{0}$, although the biomass required to support the MSY, $B_{M S Y}$, is $29 \% B_{0}$ (the long-term average biomass using a CAY strategy).

The current TACC and recent catches are at levels that are unlikely to allow the stock to move toward a level that will support the MSY. For there to be a more than $50 \%$ chance that the biomass will rebuild, the $1994-95$ reported catch would need to be no more than $7600-12500 \mathrm{t}$.

Rebuilding of the stock is likely to be slow. Even with no fishing, it is estimated that it will take until somewhere between 2000-01 and 2002-03 (depending on which data assumption is used) before there is more than a $50 \%$ chance that the biomass is greater than $B_{M S Y}$.

The best estimates of long-term yield for the Chatham Rise are that (after rebuilding) it could sustain annual removals of about 6300 t under an MCY policy, or average removals of about 8000 t under a CAY policy. These estimates are uncertain because of uncertainties in life history parameters.

### 1.2 Puysegur

There is limited information available to assess the state of the Puysegur Bank orange roughy stock. The virgin biomass, and current stock size, cannot be estimated. However, commercial catch and effort data show a marked decline in most months over the period from 1991-92 to 1993-94. These data suggest strongly that the Puysegur stock has been substantially reduced by the fishery in recent years, and the 1993-94 catch level of 5000 t is probably not sustainable.

### 1.3 Waitaki

There is currently insufficient information available to assess the state of the Waitaki orange roughy stock. Results of a trawl survey in 1992 indicate it is a small and localised spawning area, with limited commercial potential for orange roughy. Catches in recent years have been small, and principally a bycatch of fishing for oreos in the region.

### 1.4 Macquarie

There is insufficient information available to assess the state of the Macquarie Ridge orange roughy stock(s). Spawning is known to occur on two hill complexes, off the Snares and Auckland Islands, but commercial catches and catch rates have been relatively low.

## 2 Introduction

### 2.1 Overview

This document updates the stock assessment of orange roughy in Quota Management Area 3B (ORH 3B) and provides yield estimates for the Chatham Rise fishery for the 1994-95 fishing year.

This assessment incorporates new genetic data on stock structure, analyses of catch and effort (on the south and east Chatham Rise, and for the Puysegur fishery), and new survey data from the Chatham Rise.

Biomass and yield estimates for the Chatham Rise fishery are calculated using stock reduction analysis, and the likely effect of various alternative management actions is assessed using forward projections. A qualitative assessment of the Puysegur fishery is made using catch per unit effort data.

### 2.2 Description of the Fishery

Orange roughy are found in waters deeper than 750 m throughout Quota Management Area 3B. Before 1991, the main fishery was concentrated along the northern, eastern, and southeastern slopes of the Chatham Rise with more than half the catch being taken between mid June and mid August. This is the period immediately before, during, and after the spawning aggregations appear at depths of $800-1000 \mathrm{~m}$ in an arc from about 70 miles north around to east of the Chatham Islands.

Since 1991 the pattern of fishing on the Chatham Rise has changed markedly (see Sections 3.2 and 4.3.1) and fisheries have developed in the Puysegur and Waitaki areas.

Some orange roughy are caught below 1250 m , the lower depth limit for the biomass indices used in the stock assessments. However, during the period of the survey, catches below this depth are only a small percentage of the total catch, and there is evidence that, at this time, fish migrate into shallower water to spawn (Coburn \& Doonan 1994).

The fishery for orange roughy on the Puysegur Bank developed from 1991. Most fishing takes place in a small area on the western side of the Bank comprising several hills and an area of undulating slope. Catches in the region are generally mixed, with black oreo and smooth oreo as well as orange roughy. Orange roughy spawn there in winter, but to date the bulk of the fishery has taken place in the spring-summer period.

Small fisheries have also recently developed in areas near the Waitaki Canyon (1992) and south of Puysegur on the Macquarie Ridge and near the Auckland Islands (1993). Catches in these fisheries have been small.

### 2.3 Literature Review

Published material relevant to this assessment is discussed in stock assessment reports over recent years (Robertson \& Mace 1988, Robertson 1989, Francis \& Robertson 1990, 1991, Francis et al. 1992, Francis et al. 1993). Other relevant papers include an analysis of CPUE (Doonan 1991), studies on ageing (Mace et al. 1990, Fenton et al. 1991, Smith et al. in press, Francis 1995a), a description of the Chatham Rise commercial fishery (Coburn \& Doonan 1994), two FARDs on the estimation of life history parameters (Doonan 1993, 1994), and papers on the risk analysis techniques used in this assessment (Francis 1992b, Cordue \& Francis 1994).

Other recent New Zealand studies include papers on recent trawl surveys (Clark \& Tracey 1993a, 1993b; McMillan \& Hart 1994a, 1994 b, 1994c, Clark \& Thomas 1994), fecundity (Clark et al. 1994), the development of the Challenger fishery (Clark \& Tracey 1994), and indicators of biomass depletion (Francis 1995b).

Other recent Australian studies include papers on orange roughy genetics (Ward \& Elliot 1993, Smolenski et al. 1993, Elliot et al. 1994), acoustic surveys (Elliot \& Kloser 1993), modelling (Campbell et al. 1993, Smith 1993), and larval development (Jordan \& Bruce 1993).

## 3 Review of the Fishery

### 3.1 TACCs and Catches

Annual orange roughy catches in area 3B averaged just over 30000 t throughout the 1980s, and then dropped to just over 20000 t because of a substantial quota cut for 1990-91, and a smaller cut for 1992-93 (Table 1).

In recent years catches have been limited by subarea by agreement between industry and the Minister. In 1991-92, after the discovery of the Puysegur fishery, it was agreed that the catch south of $46^{\circ} \mathrm{S}$ would be at least 5000 t . In the following year (1992-93) seven subareas were defined (Francis et al. 1993, figure 0) with the following catch limits:

| North western Chatham Rise | 3500 t |
| :--- | ---: |
| Survey Box | 0 t |
| East Chatham Rise | 4500 t |
| South Chatham Rise | 6000 t |
| Waitaki | 300 t |
| Puysegur | 5000 t |
| Exploratory area | 2000 t |

These subareas are of administrative rather than scientific significance. In particular they should not be thought of as defining stocks. For example, the eastern boundary of the Waitaki subarea passes through the middle of the Waitaki fishing ground and there is solid evidence against considering the survey box and east Chatham Rise as separate (Coburn \& Doonan 1994).

In 1993, these catch limits were extended to the 1993-94 fishing year, but the Minister warned of the likelihood of further catch reductions "in the absence of a dramatic improvement in the conclusions of the stock assessment over the next year" (Anon. 1993).

Table 1: Annual reported catches and TACs of orange roughy from ORH 3B. (Catches from 1978-79 to 1985-86 are from Robertson \& Mace (1988) and from 1986-87 to 1992-93 from Fisheries Statistics Unit and Quota Monitoring System data), $\boldsymbol{\text { , no TAC }}$

| Fishing year | Reported catch (t) | TAC (t) |
| :--- | ---: | ---: |
| $1979-80 \dagger$ | 11800 | - |
| $1980-81 \dagger$ | 31100 | - |
| $1981-82 \dagger$ | 28200 | 23000 |
| $1982-83^{*}$ | 32605 | 23000 |
| $1983-84^{*}$ | 32525 | 30000 |
| $1984-85 \ddagger$ | 29340 | 30000 |
| $1985-86 \ddagger$ | 30075 | 29865 |
| $1986-87 \ddagger$ | 30689 | 38065 |
| $1987-88 \ddagger$ | 24214 | 38065 |
| $1988-89 \ddagger$ | 32785 | 38300 |
| $1989-90 \ddagger$ | 31669 | 32787 |
| $1990-91 \ddagger$ | 21521 | 23787 |
| $1991-92 \ddagger$ | 23269 | 23787 |
| $1992-93 \ddagger$ | 20077 | 21300 |

( $\dagger$ Catches for 1979-80 to 1981-82 are for a April-March fishing year.

- Catches for 1982-83 and 1983-84 are 15 month totals to accommodate the change over from an April-March fishing year to an OctoberSeptember fishing year. The TAC for the interim season, March to September 1983, was 16125 t .
$\ddagger$ Catches from 1984-85 onwards are for a October-September fishing year.)


### 3.2 Recent Changes in the Fishery

Substantial changes in the geographical distribution of the ORH 3B fishery from 1990-91 were documented by Francis et al. (1993, section 3.2.2) and are partly seen in Tables 2 and 3. The major feature of these changes was a shift of catch and effort from the Box to hills on the east Rise, and to other areas ouside the Chatham Rise. In the last two years (1991-92 and 1992-93) changes have been constrained by the subdivision of the area (see Section 3.1).

Table 2: ORH 3B catches by area from 1978-79 to 1992-93, by tonnage (to the nearest 100 t) and by percentage (to the nearest percent) of the total ORH 3B catch. All years are from 1 October. 30 September

| Year | Northwest |  | South |  | Spawning box |  | East |  | Non-Chatham |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $t$ | \% | t | \% | t | \% | $t$ | \% | $t$ | \% |
| 1978-79 | 0 | 0 | 0 | 0 | 11500 | 98 | 300 | 2 | 0 | 0 |
| 1979-80 | 1200 | 4 | 800 | 3 | 27900 | 90 | 1200 | 4 | 0 | 0 |
| 1980-81 | 8400 | 30 | 3700 | 13 | 16000 | 57 | 100 | 0 | 0 | 0 |
| 1981-82 | 7000 | 28 | 500 | 2 | 16600 | 67 | 800 | 3 | 0 | 0 |
| 1982-83 | 5400 | 35 | 4800 | 31 | 4600 | 30 | 600 | 4 | 0 | 0 |
| 1983-84 | 3300 | 13 | 5100 | 21 | 15000 | 61 | 1500 | 6 | 0 | 0 |
| 1984-85 | 1800 | 6 | 7900 | 27 | 18400 | 63 | 1100 | 4 | 0 | 0 |
| 1985-86 | 3700 | 12 | 5300 | 18 | 17000 | 56 | 4100 | 13 | 0 | 0 |
| 1986-87 | 3200 | 10 | 4900 | 16 | 20200 | 66 | 2400 | 8 | 0 | 0 |
| 1987-88 | 1600 | 7 | 6800 | 28 | 13500 | 56 | 2300 | 10 | 0 | 0 |
| 1988-89 | 3800 | 12 | 9200 | 28 | 16700 | 51 | 3100 | 9 | 0 | 0 |
| 1989-90 | 3300 | 10 | 11000 | 35 | 16200 | 51 | 1100 | 3 | 200 | 1 |
| 1990-91 | 1500 | 7 | 6900 | 32 | 6100 | 28 | 6100 | 29 | 900 | 4 |
| 1991-92 | 300 | 1 | 2200 | 9 | 1000 | 4 | 12000 | 52 | 7800 | 34 |
| 1992-93 | 4100 | 20 | 5200 | 26 | 100 | 0 | 4400 | 22 | 6300 | 32 |

Table 3: ORH 3B catches outside the Chatham Rise, by tonnage (nearest 50 t) and by percentage (nearest $1 \%$ ) of the total ORH 3B catch. For this table the areas were defined by the following rectangles: Puysegur - $46-47.5{ }^{\circ} \mathrm{S}, 165-166.5^{\circ} \mathrm{E}$; Waitaki - 45-45.5 ${ }^{\circ}$ S, 171-172 ${ }^{\circ} \mathrm{E}$; Macquarie - 475-52 ${ }^{\circ} \mathrm{S}, 163-167^{\circ} \mathrm{E}$. All years are from 1 October to 30 September

|  | Puysegur |  | Waitakl |  | Macquarie |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Season | t | \% | $t$ | \% | t | \% |
| 1989-90 | 150 | 0 | 0 | 0 | 0 | 0 |
| 1990-91 | 850 | 4 | 0 | 0 | 50 | 0 |
| 1991-92 | 6900 | 30 | 650 | 3 | 0 | 0 |
| 1992-93 | 5450 | 27 | 50 | 0 | 750 | 4 |

In 1992-93, $77 \%$ of the Chatham Rise catch was taken from three hill complexes - one in each of the three subareas where fishing was allowed: Big Chief in the South, Andes in the East, and Graveyard in the Northwest (Figure 1). Catch rates and bycatch fraction varied markedly amongst these hills, with catch rates highest (and bycatch lowest) at Graveyard, and lowest (with highest bycatch) at Big Chief (Table 4). The bycatch fraction at Big Chief has risen steadily over the five years it has been fished (Table 4). In the South, catches from the other main hills were only 1-200 t , with the oreo bycatch often being as much or more (Figure 1).

| Area Year | Catch <br> (t) | Catch rate (t per tow) | Bycatch fraction |
| :---: | :---: | :---: | :---: |
| Graveyard (42050'S - 42 ${ }^{\circ} 40^{\circ} \mathrm{S}, 179^{\circ} 53^{\circ} \mathrm{W}-179^{\circ} 48^{\circ} \mathrm{W}$ ) |  |  |  |
| 1991-92 | 100 | 2.7 | 0.35 |
| 1992-93 | 3300 | 11.0 | 0.11 |
| Andes ( $44^{\circ} 15^{\circ} \mathrm{S}-44^{\circ} 06^{\prime} \mathrm{S}, 174^{\circ} 35^{\circ} \mathrm{W}-174^{\circ} 20^{\circ} \mathrm{W}$ ) |  |  |  |
| 1991-92 | 7000 | 9.7 | 0.34 |
| 1992-93 | 2600 | 8.5 | 0.25 |
| Big Chief ( $44{ }^{\circ} 5^{\prime}-44^{\circ} 35^{\prime} \mathrm{S}, 175^{\circ} 25^{\circ} \mathrm{W}-175^{\circ} 05^{\prime} \mathrm{W}$ ) |  |  |  |
| 1988-89 | 1000 | 5.1 | 0.12 |
| 1989-90 | 2800 | 5.3 | 0.16 |
| 1990-91 | 3100 | 6.9 | 0.20 |
| 1991-92 | 800 | 6.0 | 0.43 |
| 1992-93 | 3000 | 4.7 | 0.64 |

The Chatham Rise catch was almost entirely ( $96 \%$ ) taken by five domestic vessels (in contrast to $1989-90$ where 18 vessels, including 10 foreign chartered vessels, took $91 \%$ ).


Figure 1. Tow positions, and catches of orange roughy and oreos by area, for the deepwater Chatham Rise fishery in 1992-93. Each dot represents a tow position where orange roughy or oreos were targeted or caught (no attempt has been made to remove obvious errors in tow positions). Within each circle or box, the upper figure is the orange roughy catch and the lower figure is the oreo catch (both in tonnes, as reported in tow-by-tow data).

### 3.3 Maori and Recreational Fishing

There is no known non-commercial catch of orange roughy in ORH 3B.

## 4 Research

### 4.1 Stock Structure

Allozyme data for 21 samples from 14 locations on the Chatham Rise and east coast of New Zealand (Figure 2, Table 5) were analysed to determine stock strucure for Chatham Rise orange roughy.

Table 5. Genetic data sample sizes (numbers of fish with aliele frequencies) by area, sampling time, and genetic locus. The 1982 and 1988 samples are as reported in Smith et al. (1991). Where no month is specified samples were collected during the spawning period. Data from some loci for samples before $12 / 93$ have been excluded because they are not comparable with the later data (different gels were used). Small numbers for some areas for EST-1 and IDH-1 are because processing was not complete at the time of this analysis

| Area (abbreviation) | Date | CCK-1 | EST-1 | GPI-1 | GPI-2 | IDH-1 | ]DH-2 | LDH-1 | LDH-2 MDH-1 |  | $\begin{gathered} \text { Locus } \\ \hline \text { MPI-1 PGM-2 } \end{gathered}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| Ritchie (Rch) | 3/94 | 90 |  | 99 | 99 |  | 96 | 99 | 96 | 96 | 94 | 99 |
| Wairarapa (War) | 4/94 | 93 |  | 96 | 96 |  | 95 | 85 | 96 | 96 | 94 | 96 |
| Kaikoura (Kai) | 3/94 | 87 |  | 96 | 96 |  | 94 | 96 | 92 | 94 | 92 | 96 |
| Hole (Hol) | 1/94 | 85 | 87 | 95 | 95 | 87 | 95 | 95 | 96 | 96 | 96 | 95 |
| Graveyard (Grv) | 1993 |  |  |  |  | 49 |  | 49 | 49 | 49 |  |  |
| Graveyard (Grv) | 1/94 | 87 | 99 | 92 | 92 | 99 | 95 | 92 | 97 | 97 | 95 | 92 |
| Box (Box) | 1982 |  |  | 90 | 106 |  |  |  |  |  |  | 100 |
| Box (Box) | 1988 |  |  | 48 | 48 |  |  |  |  |  |  | 48 |
| Box (Box) | 1993 |  |  |  |  | 46 |  | 47 | 47 | 47 |  |  |
| Smiths City (Smt) | 1993 |  |  |  |  |  | 48 |  | 48 | 48 | 48 |  |
| Smiths City (Smt) | 12/93 | 78 | 8 | 93 | 93 | 8 | 92 | 93 | 92 | 92 | 92 | 93 |
| Colliery (Col) | 12/93 | 40 | 97 | 97 | 40 | 96 | 97 | 97 | 97 | 97 | 97 |  |
| Possum (Pos) | 1993 |  |  |  |  |  | 47 |  | 47 | 47 | 47 |  |
| Possum (Pos) | 12/93 | 75 | 12 | 87 | 87 | 12 | 93 | 87 | 95 | 95 | 95 | 87 |
| Big Chief (BCh) | 1993 |  |  |  |  |  | 49 |  | 49 | 49 | 49 |  |
| Big Chief (BCh) | 12/93 | 76 |  | 93 | 93 |  | 91 | 93 | 93 | 93 | 93 | 93 |
| Hegerville (Heg) | 1/94 | 84 | 98 | 94 | 94 | 98 | 88 | 94 | 91 | 90 | 91 | 94 |
| Kiso (Kis) | 1/94 | 70 | 73 | 70 | 70 | 73 | 71 | 70 | 71 | 71 | 71 | 70 |
| Kiso (Kis) | 3/94 | 27 | 13 | 27 | 27 | 13 | 26 | 27 | 24 | 25 | 24 | 27 |
| SWMemoo (SWM) | 3/94 | 51 | 55 | 51 | 52 | 55 | 54 | 52 | 53 | 52 | 53 | 52 |
| Waitaki (Wai) | 3/94 | 47 | 47 | 48 | 48 | 47 | 47 | 47 | 47 | 47 | 48 | 47 |

A chi-square test on the 1993 and 1994 samples showed strong heterogeneity with most of the heterogeneity arising from two loci, $\mathrm{IDH}-2$ and MDH-1; this was also true when the test was restricted to the Chatham Rise samples (Table 6). (Alleles that were too rare to use in chi-square tests were lumped with adjacent alleles; this made 4 loci monomorphic - GPI-2, IDH-1, LDH-1, and LDH-2.)

Table 6. Results of chi-square tests for heterogeneity amongst the 1993 and 1994 samples: presented by individual locus and for all loci combined.

|  | Chatham Rise samples |  |  |  | All samples |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  | chisq | dof | $\mathbf{P}$ | chisq | dof | $\mathbf{P}$ |
| CCK-1 | 6.4 | 9 | 0.701 | 22.2 | 13 | 0.053 |
| EST-1 | 17.0 | 16 | 0.384 | 29.7 | 18 | 0.041 |
| GPI-1 | 14.3 | 18 | 0.708 | 28.5 | 26 | 0.335 |
| IDH-2 | 93.8 | 14 | 0.000 | 123.9 | 18 | 0.000 |
| MDH-1 | 37.2 | 14 | 0.001 | 156.2 | 18 | 0.000 |
| MPI-1 | 13.4 | 14 | 0.494 | 16.8 | 18 | 0.539 |
| PGM-2 | 9.0 | 9 | 0.438 | 11.5 | 13 | 0.573 |
|  |  |  |  |  |  |  |
| Combined | 191.1 | 94 | 0.000 | 388.6 | 124 | 0.000 |

To determine possible stock boundaries, a series of heterogeneity tests were done to see which groups of contiguous locations could be considered genetically homogeneous (or, rather, not significantly heterogeneous). These tests identified four possible stock boundaries: Kaikoura/Hole, Hole/Graveyard, Kiso/SWMernoo, and SWMernoo/Waitaki (Figure 3).


Figure 2. Approximate positions of 14 locations from which genetic samples were taken (see Table 5).


Figure 3. Schematic diagram illustrating the extent of genetic homogeneity between contiguous sampling locations. Vertical arrows show how far it is possible to move from the stock at the tail of the arrow before significant heterogeneity is found (solid line, $P>0.1$; broken lines, $P>0.05$ ). Horizontal lines identify possible stock boundaries suggested by these test results. Location abbreviations are given in Table 5.


Figure 4. Frequency of the most common allele at the IDH-2 and MDH-1 loci plotted by sample location (there were only two non-rare alleles at each of these loci). Where repeat samples were avallable, frequencies for individual samples are plotted as '1' and '2'; horizontal lines are as in Figure 3.
Location abbreviations are given in Table 5.

However, an examination of repeat samples at Graveyard and Kiso casts doubt on three of these possible stock boundaries. The 1993 Graveyard sample is genetically closer to that from Hole than to the 1994 sample from Graveyard, and the March 1994 sample from Kiso was closer to both SWMernoo and Waitaki than to the January Kiso sample (Table 7). This suggests that gene frequencies at a given location can vary over time to an extent that is comparable to the between-location variation that generated some of the possible boundaries in Figure 3. Chi-square tests showed marginally significant differences between repeat samples at Graveyard and at Kiso but no significant difference overall (Table 8).

| Table 7: Genetic distance boundaries in Figure 3 Graveyard 1993 |  | atrices (Nei's dista <br> Graveyard 1994 | 1975) |
| :---: | :---: | :---: | :---: |
|  |  | 92 |  |
| Graveyard 1993 | - | 80 |  |
|  | Kiso 3/94 | SWMemoo | Waitaki |
| Kiso 1/94 | 91 | 118 | 131 |
| Kiso 3/94 | - | 27 | 47 |
| SWMemoo | - | - | 73 |

Table 8. Results of chi-square tests for differences between repeat samples at the same location: presented by individual locus and

| for all loci combined |  |  |  |
| :--- | ---: | ---: | ---: |
| Area | chisq | dof | $\mathbf{p}$ |
| Graveyard | 7.2 | 3 | 0.064 |
| Smiths City | 3.1 | 3 | 0.378 |
| Possum | 2.0 | 3 | 0.573 |
| Big Chief | 1.1 | 4 | 0.893 |
| Kiso | 14.9 | 9 | 0.095 |
|  |  |  |  |
| Combined | 28.3 | 22 | 0.166 |

The reason that stock boundaries at Hole/Graveyard, Kiso/SWMernoo, and SWMernoo/Waitaki are inferred when repeat samples are lumped (Figure 3), but rejected when they are not, is illustrated in gene frequency plots (Figure 4). These plots also show a reasonably clear Kaikoura/Hole boundary. However, it would seem prudent to collect repeat samples on either side of this boundary to confirm it.

The 1982 and 1988 Box sample data were excluded from the above analyses because they were for only three loci (Table 5), and no significant heterogeneity was found at these loci in the later samples (Table 6). However, there was significant heterogeneity between the two early samples at two of these three loci (Table 9) and allele frequency differences between these samples (at the same location) were of the same degree as between-location differences in the 1993 and 1994 samples (Figure 5).

Table 9. Results of chi-square tests for differences between the 1982 and 1988 Box samples: presented by individual locus and for

| all loci combined. |  | dof | $\mathbf{P}$ |
| :--- | ---: | ---: | ---: |
| Locus | chisq | 2 | 0.026 |
| GPI-1 | 7.30 | 1 | 0.032 |
| PGM-1 | 4.63 |  | 0.008 |
|  |  | 3 |  |

Two further sources of information are relevant to the question of stock structure for ORH 3B: mitochondrial DNA analyses and fish distribution patterns.

Mitochondrial DNA analyses found no differences between the north Chatham Rise, Ritchie Bank, and Cook Canyon spawning populations. Samples from the Waitaki and Puysegur fisheries were similar to each other but different from the Chatham Rise/Ritchie/Cook samples (P.J. Smith, MAF Fisheries Greta Point, pers. comm.).

More or less simultaneous spawning has been observed at Puysegur, Cook Canyon, Waitaki, and Ritchie, and at a number of locations on the Chatham Rise - in the Box, and on hill complexes in the northwest, northeast, southeast, and southwest. The Macquarie Ridge orange roughy fishing ground is well separated from other orange roughy grounds.



Figure 5. Gene frequencies, showing sample location, for the two loci (PGM-2 and GPI-1) at which there was significant heterogeneity between the 1982 and 1988 Box samples (Table 9). PGM. 2 had only two non-rare alleles; GPI-1 had three. Location abbreviations are given in Table 5.

The abundance of orange roughy in the spawning box is low outside the spawning season, and there is uncertainty as to the origin of the fish that spawn there. However, there is evidence from analysis of commercial catch-effort data from 1979 to 1988 that fish move into and out of the spawning box from the northeast and east, and there is an indication that there may be movement from the west. This suggests that, if there is more than one stock, there could be substantial mixing between those spawning in the box and fish in the east.

On the basis of the above information it was decided to treat ORH 3B as consisting of four stocks - Chatham Rise, Waitaki, Puysegur, and Macquarie Ridge - which are separate from stocks in other QMAs.

Although there is substantial genetic heterogeneity within the Chatham Rise, there are no clear stock boundaries. The continuous distribution of orange roughy around most of the Rise, the existence of genetic differences between samples from the same location at different times, and known migration patterns within the Rise, all suggest that it will be difficult to clearly define more than one stock on the Chatham Rise. Allozyme and DNA work is continuing in an attempt to define stock relationships for orange roughy in and adjacent to ORH 3B.

### 4.2 Resource Surveys

Surveys that have been conducted in ORH 3B are shown in Table 10.

Table 10: Major research surveys for orange roughy in ORH 3B since 1982. Surveys were of two types: those aimed at estimating biomass or at locating and sampling juveniles. (N, S, and E refer to the noth, south, and east Chatham Rise, respecively.)

| Date month/year | $\begin{aligned} & \text { Area } \\ & \left(\mathrm{km}^{2}\right) \end{aligned}$ | Survey <br> Type | Vessel | Survey area | Species |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 8-9/1982 | 25000 | biomass | Kaltan | N | ORH |
| 7/1984 | 5000 | biomass | Otago Buccaneer | N | ORH |
| 7/1985 | 5000 | biomass | Otago Buccaneer | N | ORH |
| 7/1986 | 5000 | biomass | Otago Buccaneer | N | ORH |
| 11/1986 | 47100 | biomass | Arrow | S | Oreos/ORH |
| 7/1987 | 5000 | biomass | Otago Buccaneer | N\& S | ORH |
| 11/1987 | 47500 | biomass | Amaltal Explorer | S | Oteos/ORH |
| 2/1988 | na | juvenile | James Cook | N | ORH |
| 5-6/1988 | na | juvenile | James Cook | N | ORH |
| 7/1988 | 5000 | biomass | Cordella | N | ORH |
| 9/1988 | na | juvenile | James Cook | N | ORH |
| 9/1988 | 72000 | biomass | Cordella | N\& S | ORH/Oreos |
| 1/1989 | na | juvenile | James Cook | N | ORH |
| 4/1989 | na | juvenile | James Cook | N | ORH |
| 7/1989 | 25500 | biomass | Cordella | N\& S | ORH |
| 8/1989 | na | juvenile | James Cook | N | ORH |
| 10/1989 | na | juvenile | James Cook | N | ORH |
| 12/1989 | na | juvenile | James Cook | N | ORH |
| 6-8/1990 | 38600 | biomass | Cordella | N\&S\&E | ORH |
| 11/1990 | 56800 | biomass | Cordella | S\&E | Oreos/ORH |
| 7/1991 | 850 | biomass | Will Watch | Puysegur | ORH/Oreos |
| 10/1991 | 56800 | biomass | Tangaroa | S \& E | Oreos/ORH |
| 6-7/1992 | 26333 | biomass | Tangaroa | N\&S\&E | ORH |
| 7/1992 | 850 | biomass | Giljanes | Puysegur | ORH/Oreos |
| 7/1992 | 645 | biomass | Giljanes | Waitaki | ORH |
| 8-9/1992 | 6180 | biomass | Tangaroo | Puysegur | ORH/Oreos |
| 10-11/1992 | 60500 | biomass | Tangaroa | S \& E | Oreos/ORH |
| 10-11/1993 | 60500 | biomass | Tangaroa | S \& E | Oreos/ORH |
| 5-7/1994 | 35108 | biomass | Tangaroa | N \& S \& E | ORH |

A stratified random trawl survey was completed during May to July 1994. This continued coverage of the Box in July (the second survey using Tangaroa) and included most other areas of the Chatham Rise as well. Some of these had been covered in previous surveys to various degrees, but the 1994 survey involved much more rigorous sampling of several areas consisting of hill and drop-off features. These areas included both known commercial features and 'new' ones identified during bathymetric survey work carried out by the fishing industry. Hence, for those parts of the 1994 survey area outside the Box, biomass indices cannot be used until further surveys are carried out.

The 1994 survey of the Box is the ninth in a series of biomass surveys (Table 11). It was carried out following the same design as for the 1992 survey (which differed from that from previous surveys in the order of occupation of the 26 strata that make up the survey area). The biomass index from this survey is highly uncertain because:

- the coefficient of variation (c.v.) for the index is very high,
- $95 \%$ of the survey biomass is associated with just one stratum (stratum 2),
- the biomass index is reduced by $66 \%$ if a single catch (from station 399) is omitted from the biomass calculation, and
- there is a strong imbalance in the sex ratio (Table 11).

Table 11: Spawning box survey biomass indices (t) for both sexes combined (with c.v.) and by sex, percentage male (by biomass) and mean length for both sexes combined. The effect of removing one high catch rate station (station 399) from the 1994 survey is also presented. - , not estimated

| Year | Combined | c.v. (\%) | Male | Female | \% male | Mean length (cm) |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| 1984 | 164073 | 147093 | 17 | 83952 | 80120 | 51 |
| 1985 | 103367 | 15 | 73744 | 73329 | 50 | 34.72 |
| 1986 | 79388 | 16 | 55481 | 47873 | 54 | 34.78 |
| 1987 | 94583 | 15 | 41216 | 38155 | 52 | 35.67 |
| 1988 | 66955 | 25 | 53730 | 40798 | 57 | 34.59 |
| 1989 | 37780 | 18 | 32799 | 34040 | 49 | 34.84 |
| 1990 | 29011 | 19 | 20269 | 17492 | 54 | 34.35 |
| 1992 | 83169 | 34 | 8170 | 20611 | 28 | 34.95 |
| 1994 | 28263 | 67 | 11401 | 71523 | 14 | 33.87 |
| 1994 (no stn 399) | 28 | 4828 | 23435 | 17 | - |  |

The most likely explanation for the sex ratio imbalance is that it is an artefact caused by the change in survey design and the way that this relates to the (unknown) pattern of movements and catchability of male and female fish over the spawning period. There was a similar (though less extreme) imbalance in the 1992 index, when $28 \%$ of the biomass was from males. In both surveys the imbalance was observed in many catches, both small and large (Figure 6). This is in strong contrast to the earlier surveys (1984 to 1990) in which the percentage of males in the biomass ranged from $49 \%$ to $57 \%$. However, it is unlikely that the imbalance in the survey sex ratio is representative of a shift in the population sex ratio. In order for such a shift to occur from fishing, there would have had to be strong imbalances in the sex ratio of the catches over the last four years, together with a very high exploitation rate, to cause this degree of change in the sex ratio of the population spawning in the survey area. There is no evidence for an imbalance in the sex ratios of the catch in the admittedly sketchy observer database (see Section 4.3.4).

In the time series of Box surveys there is a pattern which is continued by the last two surveys. The area in which moderate to large catch rates (greater than 1 t.n.mile ${ }^{-1}$ ) were obtained has progressively shrunk (Figure 7, Table 12). In 1994 this area was confined to one stratum, which contributed $95 \%$ of the biomass index. However, the number of large catch rates (greater than 10 t.n.mile ${ }^{-1}$ ) has not shown a similar trend. Therefore, it is not possible to say whether mean fish densities within the high catch rate area have decreased or not.

Table 12: Incidence of orange roughy catch rates exceeding I t.n.mile ${ }^{-1}$ in each of the nine Box surveys. The 'No. strata' columns indicate the number of strata (or 26 ) which contained at least one tow in the indicated range of catch rates. The 'percentage' columns indicate the proportion of tows which were contained in the indicated range of catch rates

| Year | (1-10 t.n.mile ${ }^{-1}$ ) |  | ( $>10$ t.n.mile ${ }^{-1}$ ) |  |
| :---: | :---: | :---: | :---: | :---: |
|  | No. strata | Percentage | No. strata | Percentage |
| 1984 | 9 | 20 | 2 | 4 |
| 1985 | 11 | 25 | 4 | 5 |
| 1986 | 14 | 25 | 3 | 3 |
| 1987 | 8 | 20 | 0 | 0 |
| 1988 | 7 | 12 | 3 | 5 |
| 1989 | 9 | 17 | 1 | 2 |
| 1990 | 6 | 13 | 2 | 2 |
| 1992 | 3 | 6 | 1 | 2 |
| 1994 | 0 | 0 | 1 | 4 |



Figure 6. Positions of stations with extreme sex ratlos: ' 0 ' - more than $\%$ female; ' $x$ ' - more than $\%$ male (others stations shown as '. '). Dotted lines indicate stratum boundaries.

 to 10 t.n.mile ${ }^{-1}$; ' $\mathrm{O}^{\prime}$ - > 10 t.n.mile ${ }^{-1}$. Dotted lines indicate stratum boundaries.

The pattern of decrease in the moderate catch rates is consistent with a steady decrease in overall biomass, with a large proportion of the biomass located in an ever-shrinking area of high fish density. Thus it is reasonable to conclude that the biomass decline indicated by the survey indices up to 1992 is continuing, and that the 1994 biomass index is unrepresentative because of one station which by chance had a high catch rate. This pattern also explains the observation of increasing c.v. that has recently become apparent (Table 11). As the area of high fish density shrinks, the number of tows with moderate and high catch rates should fall, and the estimate of mean density in this area becomes less precise. If this hypothesis is true, the time sequence of male biomass estimates may be more representative of the trend in total biomass than the female biomass (Table 11) because the male biomass estimates do not show such a strong increase in 1994.

### 4.3 Other Studies

### 4.3.1 Catch and Effort in the South and East Chatham Rise

Catch and effort data for the south and east Chatham Rise between 1978 and 1993 (inclusive) were analysed to determine major trends in the fishery in these areas. This work extends earlier analyses for the south Rise (Francis \& Robertson 1991, Francis et al. 1992).

The area covered by these analyses is that south of $43.3^{\circ} \mathrm{S}$ or east of $175^{\circ} \mathrm{W}$ (Figure 8). The boundary at $175^{\circ} \mathrm{W}$ defines the eastern end of the survey box and has been used since October 1992 in restricting catches (see Section 3.1). However, there appears to be no natural boundary for orange roughy near this line. Catches and catch rates of orange roughy decline smoothly from west to east in this area.

In the south and east rise between 1978 and 1993 (inclusive) there were 50000 tows where either orange roughy or oreos were targeted or caught. Of these tows only 18500 targeted orange roughy but, except in the early years (before 1983), target tows accounted for most $(87 \%)$ of the orange roughy catch (Figure 9B). The number of target tows increased steadily to a peak of 2600 in 1988 and has declined steadily since (Figure 9A). Trends in orange roughy catch rates and oreo bycatches are shown in Figure 9C-F; these trends are discussed in the context of three subfisheries that are defined below.

In Figure 9, and throughout this section, catch rates are calculated as t per tow (total tonnage divided by number of tows); oreo bycatch is calculated only for tows that targeted orange roughy; and percentage oreo bycatch is calculated as 100 times the total oreo catch divided by the sum of orange roughy and oreo catches (using only those tows that targeted orange roughy). All references to years refer to calender, rather than fishing, years.

To simplify the graphical presentation of the data a "longitude" was defined for each tow on the south and east Rise as follows. The tows in this area lie along a long thin strip that runs roughly eastward from $172^{\circ} \mathrm{E}$ to $175^{\circ} \mathrm{W}$, then approximately northeast to $174^{\circ} \mathrm{W}$, and finally approximately northwest to the edge of the "Box" (Figure 8). The "longitude" for each tow describes how far it is along this strip. Specifically, each tow was projected onto the fine dotted line marked in Figure 8, and the "longitude" of each tow was set equal to the distance along this line, where the unit of distance is about 80.2 km , so that, in the horizontal portion of the line, "longitude" is the same as longitude.


Figure 8. Positions of all tows that targeted or caught either orange roughy or oreos, 1978-1993. The solid lines define the areas (Northwest, Box, East, South) to which catch restrictions have applied from October 1992. The analysis of catch and effort data described in the text for the south and east Rise covers the area which is either south of $43.3^{\circ} \mathrm{S}$ or east of $175^{\circ} \mathrm{W}$. The fine dotted Iine is the axis for "Iongitude" (see text for details); the "longitudes" of "corners" in this axis are marked ("185","187.2","188.3"). No attempt has been made to remove obvious errors in tow positions.


Figure 9. Analysis by calender year of tows in the south and east Rise (south of $43.3^{\circ} \mathrm{S}$ or east of $175^{\circ} \mathrm{W}$ ) that caught or targeted orange roughy (ORH) or oreos (OEO).


Figure 10．Maximum annual orange roughy catch by 0.1 degree＂longitude＂steps，with 13 areas of high catch Identified（see text for definition of＂longitude＂）．


Figure 12．Percentage by year of orange roughy target tows that are $\leq \mathbf{3 0} \mathbf{~ m i n}$ ．In duration for each of three sub－fisheries： south，east／flat，and east／pinn．


Figure 11．＂Longitudes＂by year for tows on the south and east Rise where more than 10 t of orange roughy was caught （see text for definition of＂longitude＂）．Vertical lines show the centres of the $\mathbf{1 3}$ areas of Figure 10.

|  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 78 |  |  |  |  |  |  |  |  |  |  |  |  |
| 79 |  |  |  |  |  |  |  |  |  |  |  |  |
| 80 |  |  |  | － |  |  |  |  | 12 | ： | ： |  |
| 81 |  |  |  |  | ： |  |  |  |  |  |  |  |
| 82 |  |  |  |  | － | 1 |  | \％ |  |  |  |  |
| 83 | ： |  |  |  |  | \％ |  |  |  |  |  |  |
| 84 |  |  |  |  |  | area | V： | 7 |  | ： | － |  |
| 85 |  |  |  |  | $\cdots$ | 县 |  | 4．${ }^{\text {\％}}$ |  |  | 1 |  |
| 86 | ． |  |  |  | 1 | Tima |  | 4， |  |  |  |  |
| 87 |  |  |  |  |  | 25 ${ }^{2}$ | － | n |  |  |  |  |
| 88 | － |  |  | ： |  | － |  | M | ： 1 |  | ． |  |
| 89 | 4 |  |  |  | i 1 | 为：${ }^{1}$ |  | 2amin | n |  | ： |  |
| 90 |  | － | － |  | ： 7 |  |  |  | ． | ， | － |  |
| 91 | ： | 2180 | 15： |  | 潞 |  | 20 | 动安 | 相 | 301840 | 1 | 34 |
| 02 | 23039 | $\pi$ | $4.5$ | 1850 | （19034 | 4rxay | 3 H | 瓦い | Exty | $7{ }^{2}$ | $\cdots$ | \％ |
| 93 |  |  | 1 | ：340． | ガt | i | i | 2 z |  | $10$ |  |  |
|  | Jan | Fob | Mar | Apr | May | Jun | Jul | Aug | Sep | Oct | Nov | Dec |

Figure 13．Dates of orange roughy target tows east of $175^{\circ} \mathrm{W}$ （＝＂longitude＂185）．Each dot represents a tow．
$84 \%$ of the total orange roughy catch from the south and east rise was caught from the 13 areas defined in Figure 10 (Table 13). Note that these areas are all east of $178^{\circ} \mathrm{E}$; the western part of the south rise (west of $178^{\circ} \mathrm{E}$ ) is primarily an oreo fishery - it accounts for only $3 \%$ of this total catch.

Table 13: Catch information for, and named features associated with, the 13 areas of high catch identified in Figure 10. Letters in parentheses, and Hills 1 to 3, refer to the features described in figure 2 of Francis et al. (1992). Percentages are calculated relative to the total south and east Rise catches

| Area | Max. ann. <br> catch (t) | \% of total <br> catch | \% of catch <br> In 1992 and 1993 | Named features |
| :---: | :---: | :---: | :---: | :--- |
| 1 | 1512 | 2.5 | 0.4 |  |
| 2 | 2456 | 5.8 | 0.1 | Trevs pinni (d) |
| 3 | 5613 | 22.9 | 1.3 | Mt. Kiso, The Busby (f,g) |
| 4 | 1111 | 2.1 | 0.5 | Amaltal pinni |
| 5 | 3633 | 6.2 | 1.0 | Hegerville (m) |
| 6 | 732 | 1.9 | 2.3 | Paranoias |
| 7 | 1785 | 2.3 | 1.3 | Condoms |
| 8 | 4947 | 11.5 | 21.1 | Cooks, Big Chief |
| 9 | 7587 | 10.8 | 50.8 | Andes, Cotopaxi, Possum |
| 10 | 771 | 1.9 | 7.1 | Not 'til Sunday |
| 11 | 1405 | 2.7 | 1.0 | (Hill 2) |
| 12 | 4132 | 7.9 | 8.2 | Smiths City (Hills 1,3) |
| 13 | 861 | 4.8 | 0.1 |  |
| Total |  | 83.5 | 95.1 |  |

The south and east Rise orange roughy fishery may be divided into three subfisheries:

```
south - west of 175.5
east/flat - primarily on the flat, east of 175 % W and before 1991
east/pinn. - primarily on pinnacles, east of 175.5}\mp@subsup{}{}{\circ}\textrm{W}\mathrm{ from }1991\mathrm{ (with some fishing
    in area 8 in 1989 and 1990) (Figure 11)
```

The eastward trend in the the south fishery (noted by Francis \& Robertson (1991)) is clearly seen in Figure 11 (also, compare columns 3 and 4 of Table 13). Another trend in this fishery is the gradual reduction in average tow duration as the ability to tow on pinnacles increased over time (Figure 12).

The contrast between the east/flat and east/pinn. fisheries is apparent in three ways. First, tows east of "longitude" 185 were not clustered around a few "longitudes" before 1991, as they were subsequently (Figure 11). Second, tows on the flat are typically longer than those on pinnacles (Figure 12). Third, the east/flat fishery was predominantly a pre- and postspawning fishery in June and August, whereas the east/pinn. fishery is prosecuted throughout the year (Figure 13).

The evolution of the three subfisheries is shown in Figure 14. For south, effort rose rapidly up to 1988 and then both catch and effort fell sharply (Figure 14A,B); catch rates peaked in 1982 at 9.2 t per tow, gradually declined through to 1989 , and have been stable at about 2.6 $t$ per tow since (Figure 14C); percentage oreo bycatch has risen consistently and is now very high, at around $60 \%$ (Figure 14E).

In the east/flat fishery, effort has been relatively low (Figure 14A), but catch rates have been high, declining from a peak of 20 t per tow in 1983 (with little effort) to 5 t per tow in 1990 (Figure 14C); oreo bycatch has been uniformly low (Figure 14D-F).


Figure 14. Analysis by calender year and subfishery of tows that caught or targeted orange roughy or oreos on the south and east Rise (south of $43.3^{\circ} \mathrm{S}$ or east of $175^{\circ} \mathrm{W}$ ). The three subfisheries are identified in the plots by the plotting symbols 's', ' P ', and ' p ' (as indicated in panel $A$ ). Catch rates in panel $C$ are for orange roughy target tows only.

Both effort and catch rose rapidly in the first three years of the east/pinn. fishery (Figure 14A,B); catch rates fluctuated about 6.6 t per tow (Figure 14C); oreo bycatch (in terms of percentage and t per tow) is not as high as in the south fishery, but has risen continuously (Figure 14E,F).

### 4.3.2 Puysegur Catch per Unit Effort

Catch and effort data from 1991-92 to 1993-94 (to June) from the central part of the Puysegur Bank fishery area have been examined. Data have been analysed by month, for two vessel categories: trawlers less than 50 m (mainly fresh fish vessels between 20 and 40 m ), and vessels greater than 50 m (mainly factory freezer trawlers of $60-70 \mathrm{~m}$ ). Months where less than 20 tows were carried out have been excluded. The average monthly catch per unit effort (total catch divided by number of tows) is shown in Table 14.

Table 14: Unstandardised CPUE (catch (t) per tow) for Puysegur Bank orange roughy, based on two size classes of vessel. -, insumficient data.

|  | $<50 \mathrm{~m}$ |  |  | $>50 \mathrm{~m}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 91-92 | 92-93 | 93-94 | 91-92 | 92-93 | 93-94 |
| Oct | - | 3.8 | 1.3 | - | 5.9 | 2.8 |
| Nov | 9.2 | 5.4 | 1.0 | 11.7 | 2.0 | 1.8 |
| Dec | 5.1 | 6.7 | - | 19.4 | 7.6 | 2.1 |
| Jan | 8.6 | - | - | 15.8 | 8.1 | 3.9 |
| Feb | 10.9 | - | - | 10.0 | 9.6 | 1.3 |
| Mar | - | - | - | - | 9.5 | 1.0 |
| Apr | - | 1.7 | - | - | 7.5 | 0.8 |
| May | - | - | - | - | - | 2.1 |
| Jun | - | - | - | 6.2 | - | 3.8 |
| Jul | - | - | - | 8.0 | - | - |
| Aug | - | - | - | 5.5 | 7.5 | - |
| Sep | - | - | - | 5.0 | 4.0 | - |

CPUE generally has shown a decline between years in all months for both vessel categories. This trend is most obvious with larger vessels, which have a more complete coverage of months, and experienced crews who have worked the grounds from the start of the fishery. CPUE over the period November - February shows a decrease from general levels of 10-20 t per tow in 1991-92, to 5-10 t per tow in 1992-93, to less than 5 t per tow in 1993-94.

### 4.3.3 Egg Surveys

In July 1994, eight vertical-haul plankton tows were carried out in the 'spawning box' to obtain fresh orange roughy eggs for use in trials of a new egg ageing technique. The method relies on being able to accurately measure the amount of DNA in each egg as an index of age. The tows targeted areas where the highest egg densities were expected. Samples between 850 m and the surface were taken using a $2 \mathrm{~m}^{2}$ ring net with 0.7 mm mesh.

Catch rates for orange roughy eggs were two orders of magnitude lower than they were in the same area 10 years earlier: maximum and mean catch rates were 0.65 and 0.33 eggs per cubic metre, compared to 57 and 15 in 1984. (These catch rates assume $100 \%$ filtration and an effective net mouth of $2 \mathrm{~m}^{2}$ for both nets; in 1984 there were 15 double oblique tows using a $2 \times 2 \mathrm{~m}$ plankton net with a 1.0 mm mesh). The comparable tows in 1992 (three oblique tows from the surface to about 780 m using multiple opening and closing nets) yielded catch rates of $0.27,0.39$, and 0.44 eggs per cubic metre - of the same order of magnitude, as the 1994 data.

The 1994 tows were notable for the number of orange roughy yolk-sac larvae caught (195 in total, and a maximum of 88 per tow). These appeared to be recently hatched as they were completely unpigmented, indicating that the eggs were hatching before they sank right to the bottom. That no larvae were found in 1984 may be because the tows were too shallow (all were shallower than 500 m ) or the mesh size of 1 mm was too large to catch yolk-sac larvae. No later stage larvae were caught in either year. No larvae were caught in 1992, even though the mesh size and depths sampled were suitable.

### 4.3.4 Sex Ratios

Chatham Rise orange roughy sex ratio data in the scientific observer database were examined to see if the sex ratio imbalance observed in the last two biomass surveys (see Section 4.2) could have been caused by fishing. Samples were taken from 843 tows between 24.11 .86 and 2.5.94, with the number of tows sampled from each area varying substantially from year to year (Table 15).

Table 15: $\quad$ Number of tows sampled for sex ratio by area and year

| Year | Northwest | Box | East | South | Total |
| :--- | ---: | ---: | ---: | ---: | ---: |
| $1986-87$ | 0 | 19 | 0 | 1 | 20 |
| $1987-88$ | 0 | 0 | 0 | 1 | 1 |
| $1988-89$ | 23 | 107 | 30 | 13 | 173 |
| $1989-90$ | 2 | 108 | 8 | 24 | 142 |
| $1990-91$ | 32 | 57 | 18 | 16 | 123 |
| $1991-92$ | 6 | 0 | 146 | 14 | 166 |
| $1992-93$ | 4 | 2 | 39 | 40 | 85 |
| $1993-94$ | 17 | 0 | 80 | 36 | 133 |

Unweighted sex ratios show that, since 1992, samples on the northwest Rise have contained mostly male fish, and that on the south Rise female fish have always dominated the samples. In the "box" and on the east Rise sex ratios have been highly variable, but approximately balanced between the sexes.

However, when the samples are weighted by catch, and the percentage of the reported catch that was sampled is taken into account, the sampling of the catch appears to have been too patchy to allow any firm conclusions to be drawn from these data.

For commercial fishing to significantly reduce the proportion of males in the population, there would have to be substantial catches of predominantly male fish. All years and areas which were well sampled and had reasonably high catches had sex ratios of about 50:50 (underlined numbers in Table 16).

| Year |  | hwest |  | Box |  | East |  | outh |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1986-87 | - | (0) | 50 | (4) | - | (0) | 33 | (<1) | 50 | (2) |
| 1987-88 | - | (0) | - | (0) | - | (0) | 43 | (<1) | 43 | (<1) |
| 1988-89 | 44 | (5) | 62 | (14) | 53 | (14) | 40 | (2) | 58 | (10) |
| 1989-90 | 54 | (<1) | 53 | (12) | 47 | (4) | 42 | (<1) | 52 | (6) |
| 1990-91 | 48 | (22) | 48 | (22) | 48 | (4) | 38 | (2) | 47 | (10) |
| 1991-92 | 70 | (10) | . | (0) | 41 | (18) | 42 | (7) | 42 | (15) |
| 1992-93 | 76 | (1) | 18 | (54) | 53 | (16) | 47 | (6) | 51 | (8) |
| 1993-94 | 64 | - | - | * | 51 | * | 45 | * | 51 | * |
| - reported catch not available for 1993-94 |  |  |  |  |  |  |  |  |  |  |

### 4.4 Biomass Estimates

### 4.4.1 Life History Parameters

Life history parameters for Chatham Rise orange roughy have been revised since the previous assessment (Table 17) on the basis of otolith ring counts from fish sampled during a survey in 1984 (Doonan 1994).

Table 17. Revised life history parameters as used in this assessment. The parameters used in the 1993 assessment are given (under
'Old') for comparison. -, not estimated

| Parameter | Symbol |  |  | New | Old |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Male | Female | Both sexes | Both sexes |
| Natural mortality | M | - | - | $0.045 \mathrm{yr}^{-1}$ | $0.04 \mathrm{yr}^{-1}$ |
| Age of recruitment | A, | 33 yr | 34 yr | - | 22.2 yr |
| Gradual recruitment | $S$ | 9 yr | 8 yr | - | 6.6 yr |
| Age at maturity | $A_{\text {m }}$ | 33 yr | 34 yr | - | 22.2 yr |
| Gradual maturity | $S_{\text {m }}$ | 9 yr | 8 yr | - | 6.6 yr |
| von Bertalanffy parameters | $L_{-}$ | 36.4 cm | 38.0 cm | - | 37.3 cm |
|  | $k$ | $0.070 \mathrm{yr}^{-1}$ | $0.061 \mathrm{yr}^{-1}$ | - | $0.07 \mathrm{yr}^{-1}$ |
|  | $t_{0}$ | -0.4 yr | -0.6 yr | - | -0.3 yr |
| Length-weight parameters | $a$ | - | - | 0.0921 | 0.0921 |
|  | $b$ | - | - | 2.71 | 2.71 |
| Recruitment variability | $\sigma_{\boldsymbol{R}}$ | - | - | 1.1 | 1.2 |
| Recruitment steepness |  | - | - | 0.75 | 0.75 |

Parameter estimates and results from ageing studies are similar to those obtained by Australian researchers working on Tasmanian stocks of orange roughy. A re-examination of the radiometric data used by Fenton et al. (1991) to age orange roughy incorporated more robust methods of analysis, but supported the conclusion that ages approaching, or exceeding, 100 years can be attained (Francis 1995a).

The new life history parameters result in a large increase in yield estimates (see Section 4.5). The reason for this can be seen by examining the effect of these changes on the dynamics of a model of a virgin population. The significant change is a substantial increase in recruitment, caused partly by the increase in $A_{r}$ and $A_{m}$ and partly by the increase in $M$, with the former change being more influential (Table 18).

Table 18: The three components of change in recruited biomass (growth, recruitment, and natural mortality - all expressed as percentages of mean recruited biomass) in a virgin orange roughy population, calculated for various sets of life history parameter values

| Life history parameters | Growth <br> $\left(\% \boldsymbol{B}_{\boldsymbol{a}}\right)$ | Recruitment <br> $\left(\% \boldsymbol{B}_{\boldsymbol{o}}\right)$ | Natural mortality <br> $\left(\% \boldsymbol{B}_{\boldsymbol{a}}\right)$ |
| :--- | ---: | ---: | ---: |
| 1993 parameters | +1.41 | +2.63 | -4.08 |
| 1994 parameters | +0.82 | +3.77 | -4.60 |
|  |  |  |  |
| 1993 parameters with $M$ as in 1994 | +1.54 | +3.02 | -4.60 |
| 1993 parameters with $A_{m}$ and $A_{\text {, }}$ as in 1994 | +0.65 | +3.39 | -4.08 |
| 1993 parameters with $S_{m}$ and $S_{\text {, }}$ as in 1994 | +1.45 | +2.59 | -4.08 |
| 1993 parameters with $\sigma_{\boldsymbol{R}}$ as in 1994 | +1.41 | +2.64 | -4.08 |
| 1993 parameters with $L_{-1} k_{n}$ and $t_{0}$ as in 1994 | +1.50 | +2.53 | -4.08 |

Recruitment is important as a determinant of yield because, as a population biomass is reduced by fishing, the growth and natural mortality terms of Table 18 are reduced proportionately, but the recruitment term does not reduce at all for the first $A_{m}$ years of the fishery, and then reduces only slowly, because of the steepness of the stock-recruit relationship.

### 4.4.2 Input Data for Biomass Estimation

The input data for the stock reduction analysis used in estimating biomass are given in Tables 11 and 19.

Table 19: Catch and overrun values used in the stock reduction analyses (a $10 \%$ overrun for 1993-94 is consistent with the assumptions made in the 1993 assessment, but these assumptions were not reviwed this year)

| Year | Catch (t) | Overrun (\%) |
| :--- | ---: | ---: |
| $1978-79$ | 11800 | 30 |
| $1979-80$ | 31100 | 30 |
| $1980-81$ | 28200 | 30 |
| $1981-82$ | 24888 | 30 |
| $1982-83$ | 15434 | 30 |
| $1983-84$ | 24818 | 30 |
| $1984-85$ | 29340 | 30 |
| $1985-86$ | 30075 | 28 |
| $1986-87$ | 30689 | 26 |
| $1987-88$ | 24214 | 24 |
| $1988-89$ | 32785 | 22 |
| $1989-90$ | 31444 | 20 |
| $1990-91$ | 20640 | 15 |
| $1991-92$ | 15529 | 10 |
| $1992-93$ | 14000 | 10 |
| $1993-94$ | $14000^{*}$ | 10 |
| * projected |  |  |

Because of uncertainties concerning the survey data (see Section 4.2), four alternative sets of data assumptions were used in estimating biomass (Table 20). The first assumption ignores the uncertainties and takes the biomass index as calculated from the entire data set. In the second assumption, the effect of removing a single very high catch rate is evaluated. Thirdly, the entire series of mean length indices was omitted because of uncertainty as to whether this contained information about recruitment. Fourthly, all data from the last two surveys (which estimated extreme sex ratios) were omitted. The first two assumptions are most comparable with the 'length, biomass' assumption in the 1993 assessment; the fourth is most comparable with the 'none' assumption.

Table 20. Four sets of data assumptions used in the stock reduction analyses. ( 83169 is the 1994 survey biomass estimate using the entire data set; $\mathbf{2 8} \mathbf{2 6 3}$ is the same estimate when one very high catch rate is omitted)

| Data | Biomass <br> indices | Mean length <br> indices | 1994 biomass <br> index (c.v., \%) |
| :--- | ---: | ---: | ---: |
| assumption | all | all | $83169(67)$ |
| 1 | all | all | $28263(45)$ |
| 2 | all | none | $28263(45)$ |
| 3 | up to 1990 | up to 1990 | - |

### 4.4.3 Chatham Rise Biomass Estimates

Biomass was estimated using the enhanced stock reduction analysis used in the 1993 assessment (Francis et al. 1993) with two modifications (see Appendix for full details). First, the $c . v$. for each biomass index was taken as the value estimated in the associated survey, rather than the median value for all indices. This was done because of the trend of increasing c.v. in these surveys (Table 11). Second, errors in the biomass indices were assumed to be lognormally, rather than normally, distributed (for an index that must be positive, normal errors are inconsistent with high c.v.s).

The effect of these changes was evaluated by repeating the 1993 analysis (for assumption 'length, biomass') with the new model assumptions. This resulted in only slight changes in the estimated biomass (Table 21).

Table 21: The effect on the 1993 stock-reduction estimates of blomass of modifications to the model assumptions (see text)

| Parameter values | Model assumptions | $\begin{array}{r} \mathbf{B}_{6} \\ (\prime 000 \mathrm{t}) \end{array}$ | $\begin{array}{r} \mathbf{B}_{19 n} \\ \left({ }^{\prime} 000 \mathrm{t}\right) \end{array}$ | $\begin{array}{r} \mathbf{B}_{1902} \\ \left({ }^{\prime} 000 \mathrm{t}\right) \end{array}$ | $\begin{gathered} \mathbf{B}_{102} \\ \left(\% \mathbf{B}_{\boldsymbol{2}}\right) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| As in 1993 | as in 1993 | 411 | 466 | 83 | 21 |
| As in 1993 | revised | 405 | 453 | 78 | 20 |

The 1994 assessment results are consistent with those from 1993. Estimates of $B_{0}$ are similar to those in the 1993 assessment (Table 22). Estimates of "current" biomass from the 1994 assessment ( $10-17 \% B_{0}$ for $B_{1994}$ ) were slightly lower than from the 1993 assessment ( $14-21 \% B_{0}$ for 1992). This decline was predicted in the 1993 assessment. (The results from the 'length only' analysis from the 1993 assessment are not included in this comparison because they were not repeated in 1994.) The value of $B_{1994}$ estimated (under the 'length, biomass' assumption) in the 1993 assessment ( $64000 \mathrm{t}, 16 \% \quad B_{0}$ ) is intermediate between the most comparable estimates (data assumptions 1 and 2 ) in the current assessment. The difference between the data assumptions has more influence on estimates of current biomass than on $B_{0}$ (Table 22, Figure 15).


Figure 15. Results from the stock reduction analysis with the four alternative data assumptions of Table 20. Arrows Indicate blomass estimates; lines represent the uncertainty about these estimates.

Table 22: Stock-reduction biomass estimates under four data assumptions (Table 20)

| Data assumption | $\begin{array}{r} B_{0} \\ (\mathbf{\prime} 000 \mathrm{t}) \end{array}$ | $\begin{array}{r} \boldsymbol{B}_{y y n} \\ \left({ }^{\prime} 000 \mathrm{t}\right) \end{array}$ | $\begin{array}{r} B_{I m} \\ (' 000 t) \end{array}$ | $\begin{gathered} B_{I S W} \\ \left(\% B_{a}\right) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: |
| 1 | 442 | 462 | 75 | 17 |
| 2 | 416 | 445 | 59 | 14 |
| 3 | 427 | 403 | 48 | 11 |
| 4 | 418 | 424 | 43 | 10 |

### 4.4.4 Biomass Estimates for Puysegur, Waitaki, and Macquarie areas

Reliable biomass estimates cannot be derived for the Puysegur stock. Trawl surveys in 1991 and 1992 were carried out with different vessels, and the time series started in 1992 with Tangaroa has not yet developed.

Similarly, trawl survey data from the Waitaki and Macquarie areas are insufficient at this stage to estimate biomass.

### 4.5 Yield Estimates

### 4.5.1 Estimation of Maximum Constant Yield (MCY)

Using the method of Francis (1992a) the maximum constant catch that can be taken indefinitely (without reducing the population below $20 \% B_{0}$ more than $10 \%$ of the time) from a population with life history parameters as in Table 17 is $1.62 \% B_{0}$. This is markedly higher than last year's figure of $1.20 \% B_{0}$ (Francis et al. 1993) because of the revised parameter values (specifically, changes in $A_{r} A_{m}$, and $M$ - see Section 4.4.1). Under continued fishing at this level the mean biomass is $50 \% \mathrm{~B}_{0}$.

The sensitivity of the MCY estimate to variations in $M$ and steepness is shown in Table 23.

Table 23: Maximum constant catch (as a percentage of $B_{0}$ ) that can be taken indefinitely (without reducing the population below $\mathbf{2 0 \%} B_{d}$ more than $\mathbf{1 0 \%}$ of the time) from an orange roughy population, calculated for different values of $M$ and steepness. -, not calculated

|  |  |  | Steepness |
| :--- | :---: | :---: | :---: |
| $M$ | 0.50 | 0.75 | 0.95 |
| 0.03 | - | 1.06 | - |
| 0.045 | 1.16 | 1.62 | 1.93 |
| 0.06 | - | 2.02 | - |

MCYs for the Chatham Rise (Table 24) were calculated following Francis (1992a). It is not yet possible to estimate MCY for the Puysegur or Waitaki populations.

Table 24. Estimates of 1994-95 beginning-of-season biomass, MCY, CAY, and MAY for the Chatham Rise population under various data assumptions (Table 20). The long-term MCY is the MCY when the biomass is greater than $20 \% B_{0}$; the MAY is the long-term average CAY. All yields are corrected for an assumed overrun of $10 \%$. All values in $t$

| Data assumption | $B_{\text {Imoses }}$ | $\mathbf{M C Y}_{\text {10x-ss }}$ | MCY $\mathbf{Y}_{\text {basem }}$ | $\mathrm{CAY}_{\text {gmas }}$ | MAY |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 80000 | 5900 | 6500 | 4800 | 8300 |
| 2 | 62000 | 4600 | 6100 | 3700 | 7800 |
| 3 | 57000 | 4300 | 6300 | 3400 | 8000 |
| 4 | 49000 | 3500 | 6200 | 2900 | 7900 |

### 4.5.2 Estimation of Current Annual Yield (CAY)

Using the method of Francis (1992a), $E_{C A Y}$, the exploitation rate producing maximum average yield (without reducing the population below $20 \% B_{0}$ more than $10 \%$ of the time) for a population with life history parameters as in Table 17 is 0.069 . This is markedly higher than last year's figure of 0.054 (Francis et al. 1993) because of the revised parameter values (specifically, changes in $A_{r}, A_{m}$, and $M$ - see Section 4.4.1). The mean catch when fishing at $E=0.069$ is $2.07 \% B_{0}$ and the mean biomass is $29 \% B_{0}$.

CAY estimates for the Chatham Rise are shown in Table 24. The sensitivity of the estimates of $E_{C A Y}$ to variations in $M$ and steepness is shown in Table 25.

Table 25: The exploitation rate, $E_{\text {cap }}$ for an orange roughy population, calculated for different values of $M$ and steepness. -, not calculated

|  |  | Steepness |  |
| :--- | :---: | :---: | :---: |
| $M$ | 0.50 | 0.75 | 0.95 |
| 0.03 | - | 0.049 | - |
| 0.045 | 0.037 | 0.069 | 0.086 |
| 0.06 | - | 0.088 | - |

It is not yet possible to estimate CAY for the Puysegur or Waitaki populations.

### 4.5.3 Sensitivity of yields to life history parameters

The effects on the estimated biomass and yield levels of changes in $M$ and the maturity/recruitment ogive are shown in Table 26. This shows that uncertainty in yields is much greater than indicated by the range of yields in Table 24. This is particularly so for the long-term yields, MCY ${ }_{\text {long-term }}$ and MAY.

Table 26: Sensitivity of estimates of 1994-95 beginning-of-season blomass, MCY, CAY, and MAY for the Chatham Rise population to changes in $M$ and the parameters that determine the maturity/recruitment ogive ( $A_{m}, A_{m} S_{m}, S_{n}$ ). The long-term MCY is the MCY when the biomass is greater than $20 \% B_{a}$; the MAY is the long-term average CAY. All yields are corrected for an assumed overrun of $\mathbf{1 0 \%}$
Data
assumption
2 $\quad \boldsymbol{M}$

### 4.6 Decision Analysis for the Chatham Rise Fishery

Forward simulations were carried over the period 1994-95 to 2001-02 for the constant-catch options shown in Table 27. The time period and catch options were chosen so that the results of the simulations should be as comparable as possible with those from the 1993 assessment. Option 6, an immediate cessation of fishing, is included to illustrate how the population might behave in the absence of fishing.

Table 27: Annual Chatham Rise catches ( $t$ ) for six constant-catch options used in forward simulations over the period 1994-95 to 2001-02, inclusive. (In the forward simulations, actual catches were assumed to be $\mathbf{1 0 \%}$ higher than the values below)

| Option | 1 | 2 | 3 | 4 | 5 | 6 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Catch (t) | 9000 | 7500 | 6000 | 5000 | 4000 | 0 |

For each catch-reduction option, four measures of fishery performance were calculated (Table 28). All four measures relate to the mid-year biomass in 2001-02: thus $P$ (Improve by 2002) is the probability that this is either greater than the "current" mid-year biomass or greater than $30 \% B_{0}$; and $\mathrm{P}\left(B_{2002}>20 \% B_{0}\right)$ and $\mathrm{P}\left(B_{2002}>30 \% B_{0}\right)$ are the probabilities that it exceeds $20 \% B_{0}$ and $30 \% B_{0}$, respectively. To enhance comparability with the 1993 assessment, two
versions of $P$ (Improve by 2002) were calculated. These are distinguished by the subscripts " 93 " and " 94 ", which denote the year which is treated as "current" in the definition just given. The measure P (No collapse) is not presented because its estimated values were uniformly high and so show very little contrast between the catch options and/or data assumptions. The measures $\mathrm{P}_{93}$ (Improve by 2002) and $\mathrm{P}\left(B_{2002}>20 \% B_{0}\right)$ are directly comparable with the similarly named measures in the 1993 assessment (after allowance for differences in the total catches over the 1993-94 to 2001-02 period). Note that, in terms of these fishery performance indicators, only the total catch over the period considered is important; the sequence of individual catches has little effect.

Table 28: Estimates of fishery performance for six catch options (Table 27), calculated for four alternative data assumptions (see Table 20)

| Data assumption | Performance measure | Option |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1 | 2 | 3 | 4 | 5 | 6 |
| 1 | $\mathrm{P}_{94}$ (Improve by 2002) | 0.49 | 0.71 | 0.89 | 0.96 | 0.99 | 1.00 |
|  | $\mathrm{P}_{93}$ (Improve by 2002) | 0.39 | 0.57 | 0.78 | 0.88 | 0.95 | 1.00 |
|  | $\mathrm{P}\left(\mathrm{B}_{2002}>20 \% B_{0}\right)$ | 0.34 | 0.45 | 0.58 | 0.67 | 0.76 | 0.97 |
|  | $\mathrm{P}\left(B_{2002}>30 \% B_{0}\right)$ | 0.09 | 0.12 | 0.18 | 0.23 | 0.28 | 0.60 |
| 2 | $\mathrm{P}_{91}$ (Improve by 2002) | 0.38 | 0.61 | 0.85 | 0.94 | 0.98 | 1.00 |
|  | $\mathrm{P}_{93}$ (Improve by 2002) | 0.28 | 0.45 | 0.69 | 0.83 | 0.92 | 1.00 |
|  | $\mathrm{P}\left(B_{2002}>20 \% B_{0}\right)$ | 0.16 | 0.26 | 0.38 | 0.48 | 0.59 | 0.93 |
|  | $\mathrm{P}\left(B_{2002}>30 \% B_{0}\right)$ | 0.03 | 0.05 | 0.07 | 0.10 | 0.15 | 0.43 |
| 3 | $\mathrm{P}_{94}$ (Improve by 2002) | 0.82 | 0.92 | 0.98 | 0.99 | 1.00 | 1.00 |
|  | $\mathrm{P}_{93}$ (Improve by 2002) | 0.77 | 0.88 | 0.95 | 0.98 | 0.99 | 1.00 |
|  | $\mathrm{P}\left(B_{2002}>20 \% B_{0}\right)$ | 0.37 | 0.49 | 0.62 | 0.70 | 0.78 | 0.98 |
|  | $\mathrm{P}\left(B_{2002}>30 \% B_{0}\right)$ | 0.10 | 0.14 | 0.20 | 0.25 | 0.32 | 0.64 |
| 4 | $\mathrm{P}_{94}$ (Improve by 2002) | 0.67 | 0.83 | 0.95 | 0.99 | 1.00 | 1.00 |
|  | $\mathrm{P}_{93}$ (Improve by 2002) | 0.58 | 0.75 | 0.89 | 0.95 | 0.98 | 1.00 |
|  | $\mathrm{P}\left(B_{2002}>20 \% B_{0}\right)$ | 0.21 | 0.29 | 0.40 | 0.48 | 0.57 | 0.89 |
|  | $\mathrm{P}\left(B_{2002}>30 \% B_{0}\right)$ | 0.06 | 0.09 | 0.11 | 0.15 | 0.19 | 0.44 |

Estimates of $\mathrm{P}_{93}$ (Improve by 2002) and $\mathrm{P}\left(B_{2002}>20 \% B_{0}\right)$ are markedly greater than the comparable values in the 1993 assessment (Figure 16). Two factors are responsible for this increase. The first is the increase in estimated productivity of orange roughy caused by the revised values of the biological parameters $A_{r}, A_{m}$, and $M$ (see Section 4.4.1). The second reason applies mainly to data assumption 3 (and to a lesser extent to assumption 4). When mean length data are not used in the assessment, estimates of recent recruitment are higher, which leads to predictions of faster rebuilding.

That estimates of $\mathrm{P}_{94}$ (Improve by 2002) are always higher than $P_{93}$ (Improve by 2002) shows that, for all data assumptions, the current assessment indicates a drop in


Figure 16. Comparison of the 1993 and 1994 estimates of the performance measures $P_{93}$ (Improve by 2002) (upper panel) and $\mathrm{P}\left(B_{2002}>\mathbf{2 0 \%} B_{a}\right)$ (lower panel). Plotting symbols ' 1 ', '2', '3', '4' correspond to data assumptions 1 to 4 in the current assessment (Table 20); symbols ' $B$ ', ' N ', and ' $L$ ' correspond to the 'Length, biomass', 'None', and 'Length' data assumptions of the 1993 assessment.
biomass between 1992-93 and 1993-94. This means that current catch exceeds estimated surplus production (otherwise the biomass would have increased).

Two performance criteria that are suggested by the performance measures given here are:
'safety' criterion: $\mathrm{P}\left(B_{2002}>20 \% B_{0}\right)>0.9$
'target' criterion: $\mathrm{P}\left(B_{2002}>30 \% B_{0}\right)>0.5$
These are analogous to the criteria used in the 1994 ORH 2A, 2B, 3A assessment (Field et al. 1994).

None of the non-zero catch options of Table 27 meets both criteria (using any of the four data assumptions). Estimates of the maximum total catch (excluding overrun) over the 8 year period 1994-95 to 2001-02 that would meet both criteria range from 0 to 14000 t (Table 29). Even with no fishing, neither of these criteria is expected to be met before 2000-01 (Table 30)

Table 29: Maximum total catch (t, excluding overrun), over the elght year period 1994-95 to 2001-02, that would meet the 'safety' or 'target' performance criteria (or both), for each of the data assumptions of Table 20

| Data |  |  |  |
| :--- | ---: | ---: | ---: |
| assumption |  |  | Performance criterion |
|  | 'safety' | 'safety' + 'target' |  |
| 1 | 13000 |  |  |
| 2 | 4000 | 9000 | 9000 |
| 3 | 17000 | 0 | 0 |
| 4 | 0 | 14000 | 14000 |
|  |  | 0 | 0 |

Table 30: Estimates of the fishing year in which the 'safety' and 'target' performance criteria are first expected to be met, assuming no fishing in the future (calculated for each of the data assumptions of Table 20)

| Data |  |  |  |
| :--- | :---: | :---: | :---: |
| assumption |  |  | Performance criterion |
|  | 'safety' | 'target' |  |
| 1 | $2000-01$ | $2001-02$ |  |
| 2 | $2001-02$ | $2002-03$ | $2001-02$ |
| 3 | $2000-01$ | $2000-01$ | $2002-03$ |
| 4 | $2002-03$ | $2002-03$ | $2000-01$ |
|  |  |  | $2002-03$ |

Estimates of the total catch which can be taken in the eight year period 1994-95 to 2001-02, and which will result in a $50 \%$ probability that the stock size in 2001-02 will be an 'improvement' on that in 1992-93, range from 58000 to 96000 t (Table 31). The 1993-94 catch of 14000 t must be added to these figures to make them comparable to the value resulting from the 1993 stock assessment (about 45000 t ). The increase in these catch levels is caused mainly by the increase in the estimated productivity of orange roughy (see Section 4.4.1.

Table 31. Estimates of total catch (excluding overrun) for the eight year period 1994-95 to 2001-02 for which $\mathrm{P}_{93}$ (Improve by 2002) equals 50\%

| Data Assumption | Total Catch (t) |
| :--- | ---: |
| 1 | 65000 |
| 2 | 58000 |
| 3 | 96000 |
| 4 | 78000 |

### 4.7 Effect of 1994-95 catch

The effect of the 1994-95 catch may be summarised by the probability that this catch will allow the biomass to 'rebuild'. In this context the term 'rebuild' means that EITHER the 1994-95 end-of-year biomass was greater than that for 1993-94 OR the 1994-95 mid-season biomass was greater than $20 \% B_{0}$. The maximum 1994-95 catch that allows at least a $50 \%$ chance that the biomass will rebuild, ranges from 7600 to 12500 t , according to the assumptions made (Table 32). These $50 \%$ rebuild catches are substantially higher than those estimated for 1993-94 in the 1993 stock assessment. The increase in these catch levels is caused mainly by the increase in the estimated productivity of orange roughy (see Section 4.1.1).

Table 32: The probability that the biomass will rebuild with various 1994-95 catch levels, and the catch that alfows a $50 \%$ chance of rebuilding, calculated for four alternative data assumptions (Table 20). ( $\mathbf{1 0 \%}$ catch overrun was assumed for 1994-95). -, not estimated

| Data |  |  |  | 1994-95 catch (t) |  | 50\% rebuild |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| assumption | $\mathbf{1 2 0 0 0}$ | $\mathbf{9 0 0 0}$ | $\mathbf{7 5 0 0}$ | $\mathbf{6 0 0 0}$ | $\mathbf{5 0 0 0}$ | $\mathbf{4 0 0 0}$ | catch (t) |
| 1 | - | 0.36 | 0.64 | 0.89 | 0.97 | 0.99 | 8300 |
| 2 | - | 0.24 | 0.52 | 0.83 | 0.95 | 0.99 | 7600 |
| 3 | 0.55 | 0.85 | 0.94 | 0.98 | 1.00 | 1.00 | 12500 |
| 4 | - | 0.62 | 0.84 | 0.96 | 0.99 | 1.00 | 9700 |

### 4.8 Sensitivity to Life History Parameters

Estimates of the performance measures of Table 28, and the $50 \%$ rebuild catches of Table 32, are rather sensitive to changes in key life history parameters (Tables 33 and 34). Thus the degree of uncertainty in these estimates is substantially greater than shown in Tables 28 and 32 (and, by implication, Table 31).

Table 33: Sensitivity of estimates of fishery performance for catch option 2 (Table 27) to changes in $M$ and the parameters that determine the maturity/recruitment ogive ( $\left.A_{m}, A_{m}, S_{m}, S_{N}\right)$, calculated for two alternative data assumptions (Table 20)

| Data assumption | M | Maturity ogive | $\mathbf{P}_{\mathbf{m}}$ (Improve by 2002) | $\mathrm{P}_{83}($ Improve by 2002) | $\mathrm{P}\left(\mathrm{B}_{2002} \mathbf{2 0 \%} \mathrm{Ba}_{\mathrm{a}}\right)$ | $\mathrm{P}\left(\mathrm{B}_{202}>\mathbf{2 0 \%} \mathrm{Ba}_{\boldsymbol{a}}\right)$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | 0.045 | 1994 version | 0.61 | 0.45 | 0.26 | 0.05 |
|  | 0.03 | 1994 version | 0.25 | 0.13 | 0.07 | 0.01 |
|  | 0.06 | 1994 version | 0.82 | 0.71 | 0.46 | 0.13 |
|  | 0.045 | 1993 version | 0.19 | 0.11 | 0.17 | 0.03 |
| 3 | 0.045 | 1994 version | 0.92 | 0.88 | 0.49 | 0.14 |
|  | 0.03 | 1994 version | 0.68 | 0.58 | 0.19 | 0.04 |
|  | 0.06 | 1994 version | 0.98 | 0.97 | 0.71 | 0.30 |
|  | 0.045 | 1993 version | 0.81 | 0.74 | 0.34 | 0.11 |

Table 34: Sensitivity of estimates of the 1994-95 ' $50 \%$ rebulid' catch to changes in $M$ and the parameters that determine the maturity/recruitment ogive ( $A_{m}, A_{m} S_{m}, S_{p}$ ), calculated for two alternative data assumptions (Table 20)
$\left.\begin{array}{lllr}\begin{array}{lll}\text { Data } \\ \text { assumption } \\ 2\end{array} & M & \begin{array}{l}\text { Maturity } \\ \text { ogive }\end{array} & 50 \% \text { rebuild } \\ \text { catch }(\mathbf{t})\end{array}\right)$

## 5 Management Implications

### 5.1 Chatham Rise Fishery

This stock assessment is based on the results of applying the stock reduction model using the following assumptions: a) there is a single stock for the entire Chatham Rise; b) the stock is accurately indexed by the time series of surveys in the Box; and c) the ranges used for the biological parameters encompass their true values. This leads to the following five conclusions:

- The mid-season 1993-94 biomass estimates ranged from 10 to $17 \% B_{0}$. These estimates are substantially less than $B_{M S Y}\left(29 \% B_{0}\right)$.
- In terms of long-term yield and rebuilding potential, the results of this assessment are more optimistic than those of last year. This is primarily because revised values of some biological parameters ( $A_{m}, A_{r}, M$ ) have caused an increase in the estimated productivity of orange roughy. However, current catch levels are about twice the estimated long-term yields ( $\mathrm{MCY}_{\text {long-term }}$ and MAY - see Table 24). There is no significant change in the estimated size of the virgin stock ( $B_{0}$ ).
- The estimated biomass continues to decline. This indicates that the current catch exceeds estimated surplus production.
- Current estimates of the catch that will result in either the 1994-95 end-of-year biomass being greater than that for 1993-94 or the 1994-95 mid-season biomass being greater than $20 \% B_{0}$ range from 7600 to $12500 t$ (excluding overrun) for 1994-95. Estimates of the maximum total catch (excluding overrun) over the period 1994-95 to 2001-02 (inclusive) that will allow a $50 \%$ chance that the biomass in 2001-02 will be an 'improvement' on that in 1992-93, range from 58000 to 96000 t. Note that catch levels that allow only a $50 \%$ probability of rebuilding are, in effect, "status quo" catches in that there are equal probabilities that the stock will rebuild or be depleted as a result of these catches.
- The low productivity of orange roughy makes rebuilding rates slow. To achieve rebuilding to $B_{M S Y}$ by 2001-02 would require a substantial reduction in catch, and might not be possible even if fishing ceased (see Table 29).

This section concludes with a discussion of the main sources of uncertainty associated with this assessment (life history parameters, stock definition, interpretation of the 1992 and 1994 survey results, recruitment, and representativeness of the survey indices). The first two are the most important. None of these sources of uncertainty give reason to believe that the status of Chatham Rise orange roughy is likely to be either more pessimistic or more optimistic than indicated by the current assessment.

## - Life history parameters

The main uncertainties concerning the biological parameters are age estimation and the steepness of the stock-recruit relationship. The former is fundamental to the determination of the productivity of orange roughy (through parameters like $M, A_{m}$, and $A_{r}$ ). Although current estimates of these parameters are more consistent than those used in 1993 with the data now available, they are still highly uncertain. This leads to similar uncertainty in estimates of yield, rebuilding catches, and fishery performance. The steepness parameter is an important component in determining longterm productivity. No data are available to check the assumed value of this parameter.

- Stock definition

Recent genetic data suggest that Wairarapa and Kaikoura fish are separate from those on the Chatham Rise, but have not resolved the issue of stock structure within the Rise. Heterogeneity within the data suggests more than one stock (as may the widespread spawning activity outside the Box), but there are no clear stock boundaries. If there is more than one stock, then (a) the size of the stock indexed by the Box surveys has been substantially reduced, but the exact level of depletion is unknown, and (b) the status of the stock(s) not indexed by the Box surveys is unknown.

## - $\quad$ Survey results

High coefficients of variation (c.v.s) and an imbalance in the sex ratios make interpretation of the 1992 and 1994 survey results difficult and uncertain. However, all four data assumptions (see Table 20) considered in this assessment lead to the conclusion that the biomass of this stock has continued to decline since the last assessment. The effect on yield estimates of this uncertainty is less than that from the life history parameters.

- Recruitment

Uncertainty with regard to recruitment centres on the interpretation of the lack of consistent change in mean length, combined with a strong decline in the biomass index. The traditional interpretation is that recent recruitment has been poor. However, the observation that mean length has not changed in other orange roughy populations
casts doubt on this interpretation, so recent recruitment may be higher than estimated for data assumptions 1 and 2 (see Table 20). The uncertainty with regard to recruitment is responsible for the differences in the assessment results based on data assumptions 1 and 2 compared to those based on assumptions 3 and 4.

## - Representativeness of survey indices

Another uncertainty is that of representativeness, i.e., the assumption that the survey data index the biomass and mean length of a single recruited Chatham Rise population of orange roughy. The main argument in favour of this assumption is that the Box covers what appears to have been, for at least most of the history of the fishery, the major orange roughy spawning ground on the Chatham Rise. A possible alternative hypothesis is that the proportion of Chatham Rise orange roughy that spawn in the Box has declined over the years. This would produce a more optimistic assessment. However, to make a major difference to the assessment the decline in the proportion spawning would need to have been considerable before the 1991 season (because the analysis with data assumption 4 uses only survey data from before 1991).

### 5.2 Puysegur Fishery

There is little information available to assess the state of the Puysegur Bank orange roughy stock. CPUE data suggest strongly that the Puysegur stock has been reduced by the fishery in the last two years. The current catch level of 5000 t is probably not sustainable. For it to be sustainable, a virgin biomass of around $300000 t$ (about $70 \%$ of the size estimated from the Box surveys on the north Chatham Rise) would be required. To reduce risks to the fishery, lower catch levels are required.

An additional research survey of the Puysegur stock was carried out in September-October 1994, after this assessment.

### 5.3 Waitaki Fishery

There has been only one research survey in this area. Results had a low level of precision, and are of limited use in assessing the state of the stock. However, the patchiness of the distribution of orange roughy was more extreme than in the other spawning grounds in New Zealand, and the very low 'background' catch rates in surrounding strata were also unusual in comparison with other fishing areas. Survey results indicated a small and localised spawning area, with limited potential to support a major target fishery for orange roughy.

### 5.4 Macquarie Ridge Fishery

Two hill complexes fished commercially in 1993 may have some potential. Spawning occurs in both areas in July. There are no quantitative data on abundance, but indications are that the distribution of orange roughy is very localised on these features.

## Acknowledgments

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## Appendix: Stock Reduction and Decision Analysis

This appendix describes the versions of stock reduction and decision analysis used in this assessment. The techniques are similar in principle to those described by Francis (1990, 1992b), but have been modified in a number of ways (including the addition of stochastic recruitment, recruitment/maturity ogives, importance sampling, and lognormal error structures).

## Overview

I first give an overview of the two techniques. Details of the methods are given in subsequent sections.

The aim of stock reduction analysis is to estimate past and present biomass for a fishery. Three types of input data are required: biological parameters, abundance (and possibly mean length) indices, and a complete catch history (Table A1). For decision analysis, we extend the catch history into the future by adding proposed future catch levels, and evaluate the likely response of the fishery to these catch levels by estimating various measures of risk or fishery performance.

Table A1. Data required, and corresponding model notation, for stock reduction and decision analyses

## Parameter Description

Subscripts
$s \quad$ sex $(1=$ male, $2=$ female $)$
$y \quad$ year
$j \quad$ indexes the $j$ th series of biomass (or mean length) indices
Biological parameters
$M_{s} \quad$ instantaneous natural mortality $\left(y^{-1}\right)$
$A_{r s} \quad$ age at $50 \%$ recruitment (y)
$S_{r s} \quad$ recruitment ogive width parameter (y)
$A_{m s}{ }^{*} \quad$ age at $50 \%$ maturity (y)
$S_{m s}^{*} \quad$ maturity ogive width parameter (y)
$L_{\text {ss }}$
$\left\{\begin{array}{c}(\mathrm{cm}) \\ \text { von Bertalanffy parameters }\left(y^{-1}\right)\end{array}\right.$
$t_{0}$
$a_{s} \quad$ length-weight
$b_{s} \quad$ (parameters $\left[W=a L^{b}, L\right.$ in $\mathrm{cm}, W$ in g$]$
$\sigma_{R} \quad$ recruitment variability
$h \quad$ stock-recruit steepness
$A_{\text {max }} \quad$ maximum age in model (y) (there is a plus group at this age)
$F_{\text {max }} \quad$ maximum possible exploitation rate
Biomass and/or mean length indices

| $O_{y j}$ | value of $j$ th index in year $y$ |
| :--- | :--- |
| $c_{y j}$ | coefficient of variation of $O_{y j}$ |

Catch history
$C_{y} \quad$ catch (historical or projected) in year $y$

[^0]To run the population model underlying these analyses we need values for the biological parameters in Table A1, and for $B_{0}$ (mean virgin biomass). In addition, the model generates a vector, $\varepsilon$, of recruitment deviates, which, together with a stock-recruit relationship, allows it to calculate the number of fish at age 1 in any year (see below).

The model is used for stock reduction and decision analyses in the following way.

## 1. Choose a trial $B_{0}$ value.

2. Generate a vector, $\varepsilon$, of recruitment deviates.
3. Run the population model to calculate the biomass and mean length in each year (together with any other quantities needed for performance or risk measures - e.g., whether it would be possible to catch the projected catches).
4. Calculate the likelihood associated with ( $\left.B_{0}, \varepsilon\right)$ and the indices, $O_{y j}$ (see below).
5. Repeat steps 2-4 with a large number ( 20000 in this assessment) of different recruitment deviate vectors.
6. Repeat steps $1-5$ with an equally spaced sequence of trial $B_{0}$ values, using the same set of recruitment deviate vectors for each trial $B_{0}$.
7. From the resulting set of model runs (in this assessment there were 380000 runs: 19 trial $B_{0}$ values times 20000 recruitment deviate vectors) select a sample of size 5000 , with replacement, with the probability of selection for any one run being proportional to the likelihood. This is called an importance sample.
8. Calculate mean biomass (or length) for any year as the mean of the biomass values calculated at step 3, where the mean is taken over all runs in the importance sample (with each run having equal weight). Use a similar procedure to calculate risk/performance measures (see below).

The importance sample is used to cut down the computation time. This is possible because the likelihood associated with each model run depends only on the historical catches, but decision analyses are typically carried out for a series of catch "histories" that are identical for the historical catches but contain different future catch levels (associated with alternative management strategies). Thus, while 380000 model runs are required for one catch history, only 5000 model runs are required for subsequent catch histories. (In fact fewer than 5000 runs are required because the importance sample is selected with replacement, so that some runs will be selected many times).

For each trial $B_{0}$ value we can calculate the sum of the associated 20000 likelihoods. The range of trial $B_{0}$ values selected at step 6 must be wide enough so that the value of this sum for the maximum and minimum $B_{0}$ values is a small fraction ( $<1 \%$, say) of the maximum value of this sum.

## The Population Model

Given input data (Table A1) and a value for $B_{0}$, the population model first calculates values for the derived parameters (Table A2) as follows.
(A1) $A_{r, l_{0}}=\operatorname{int}\left(A_{r m}-S_{r m}\right)$
(A2) $A_{r, h i}=\operatorname{int}\left(A_{r s}+S_{r s}+0.999\right)$
(where int $(x)$ returns the greatest integer less than or equal to $x$ )
(A3)

$$
r_{i s}=\left\{\begin{array}{cc}
0 & 1 \leq i<A_{r, l o} \\
1 /\left[1+19^{\left.\left(A_{m}-i\right) / s_{m}\right]}\right. & A_{r, l o} \leq i \leq A_{r, h i} \\
1 & A_{r, h i}<i \leq A_{\max }
\end{array}\right.
$$

(A4) $p_{i s}=\frac{1-r_{i s}}{1-r_{i-1, ~}}$
(A5) $\quad L_{i s}=L_{c_{d}}\left[1-e^{k_{d}\left(i-t_{0}\right)}\right]$

$$
\text { (AO) } \quad w_{i s}=10^{6} a_{\mathrm{f}} L_{i t}^{b_{t}}
$$

(A7)

$$
\text { (A8) } \quad \overline{R_{0}}=\frac{2 B_{0}}{\Sigma_{s} \theta_{s}}
$$

(A9) $\quad \alpha=0.5 \theta_{2}\left(1-\frac{h-0.2}{0.8 h}\right)$
(A10) $\quad \beta=\frac{h-0.2}{0.8 h \overline{R_{0}}}$

Note the distinction between two uses of "recruitment": recruitment to the fishery, which happens at maturity (and involves parameters $A_{r s}, S_{r s}, N_{y i s}, r_{i s}, p_{i s}, A_{r, j}, A_{r n i}$ ), and recruitment to the population, which happens at age 1 (and involves parameters $\sigma_{R}, h, \bar{R}_{y}, R_{y}, \varepsilon_{y}$ ).

Table A2. Description and notation for model parameters and variables (in addition to those in Table A1). These are given in the order in which they are introduced in the text.
Parameter Description
Subscripts
$s \quad$ sex $(1=$ male, $2=$ female $)$
$i \quad$ age (y)
$y \quad$ year
$j \quad$ indexes the $j$ th series of biomass (or mean length) indices
Derived parameters
$\begin{array}{ll}A_{r, t o} & \text { (range of ages for which there is } \\ A_{r, h i} & \text { lartial recruitment/maturity }\end{array}$
$r_{i s} \quad$ proportion of fish (of age $i$ and sex $s$ ) that are recruited (= mature) in the virgin
population
$p_{\text {is }} \quad$ proportion (of fish of age $i-1$ and sex $s$ that were unrecruited in one year) that will still be unrecruited in the next year
$L_{i s} \quad$ mean length of fish of age $i$ and sex $s$
$w_{i s} \quad$ mean weight of fish of age $i$ and sex $s$
$\theta_{s} \quad$ constant used in calculating $R_{0}$
$\bar{R}_{0} \quad$ mean virgin recruitment (at age 1)
$\alpha \quad$ (Beverton \& Holt stock-recruit
$\beta \quad$ lrelationship parameters
Variables calculated at each model iteration

| $\bar{R}_{y}$ | $\quad$expected recruitment (at age 1) in year $y$ <br> $N_{y i s}$ |
| :--- | :--- |
| number of fish (age $=i$, sex $=s)$ in year $y$ |  |
| $\varepsilon_{y}$ | recruitment deviate for cohort that recruits (at age 1) in year $y$ |
| $R_{y}$ | actual recruitment (at age 1) in year $y$ |

The model then initialises the population structure by generating a vector of standard normal variates $\left\{\varepsilon_{1, i}: 1 \leq i \leq A_{\text {max }}\right\}$ and calculating numbers at age as follows.

$$
\begin{aligned}
& \text { (A11) } \quad R_{1-i}=\bar{R}_{0} e^{\epsilon_{1-i}-\sigma_{R}-0.50_{R}^{2}} \quad \text { for } 1 \leq i \leq A_{\max } \\
& \text { (A12) } \quad N_{0<s}=\left\{\begin{array}{cl}
0.5 R_{1-i} e^{-M_{\Omega}(i-1)} & 1 \leq i<A_{\max } \\
0.5 \overline{R_{0}} \frac{e^{-M_{,}\left(A_{\max }-1\right)}}{1-e^{-M_{s}}} & i=A_{\max }
\end{array}\right.
\end{aligned}
$$

$$
(A 13) \quad N_{0 i s}^{r}=N_{0 \leq s} r_{i s}
$$

$$
\text { (A14) } \quad N_{0 i s}^{u}=N_{0 i s}\left(1-r_{i s}\right)
$$

An iterative procedure is used to calculate numbers at age for each successive year. At each step, a standard normal variate, $\varepsilon_{y}$, is generated, and the following equations are evaluated.

$$
\begin{gathered}
\text { (A15) } \overline{R_{y}}=\frac{B_{y-1,2}^{3}}{\alpha+\beta B_{y-1,2}^{3}} \\
\text { (A16) } \quad R_{y}=\bar{R}_{y} e^{\epsilon, 0_{R}-0.50_{R}^{2}} \\
\text { (A17) } \quad N_{y t s}^{u}=\left\{\begin{array}{rl}
N_{y-1, i-1, s}^{u} e^{-M_{s}} & 2 \leq i<A_{r, l} \\
N_{y-1, l-1, s}^{u} p_{i s} e^{-M_{s}} & A_{r, l o} \leq i \leq A_{r, k} \\
0 & A_{r, k i}<i \leq A_{\max }
\end{array}\right.
\end{gathered}
$$



The biomasses, exploitation rate, and mean length associated with each year are calculated using the following equations.
(A19) $\quad B_{y s}^{1}=\Sigma N_{y i s}^{r} w_{i s}$
(A20) $\quad B_{y s}^{2}=B_{y s}^{1} e^{-M_{z}}$
(A21) $\quad B_{y s}^{3}=B_{y s}^{2}\left(1-0.5 F_{y}\right)$
(A22) $\quad B_{y s}^{4}=B_{y s}^{2}\left(1-F_{y}\right)$
(A23) $\quad F_{y}= \begin{cases}\frac{C_{y}}{\Sigma_{s} B_{y s}^{2}} & \text { if } \frac{C_{y}}{\Sigma_{s} B_{y s}^{2}}<F_{\text {max }} \\ F_{\max } & \text { otherwise }\end{cases}$
(A24) $\overline{L_{y}}=\frac{\sum_{i v} N_{y t s}^{r} L_{i s}}{\Sigma_{i t} N_{y i s}^{r}}$

## Calculation of Likelihood

In this section I show how the likelihood associated with $\left(B_{0}, \varepsilon\right)$ and the indices $O_{y j}$ is calculated.

For all indices, it is assumed that $O_{y j}$ is either normally (for mean length) or lognormally (for biomass) distributed with mean $q_{j} E_{y j}$ and coefficient of variation $c_{y j}$. The $c_{y j}$ are assumed known, the $q_{j}$ is to be estimated, and the $E_{y j}$ are calculated from the population model - either as mid-season biomass $\left(B_{y l}^{3}+B_{y 2}^{3}\right)$ or as $L_{\text {. }}$. (Note that, because we are more interested in the trend in observed mean lengths than their absolute values, and because the $L_{\text {ss }}$ are not well determined, observed mean lengths are treated as relative, rather than absolute, indices.)

If the recruitment deviates, $\varepsilon$, are such that the historical catches could not have been caught, the likelihood, $L$, is set to zero. Otherwise, it is calculated as

$$
\begin{equation*}
L=e^{\sum_{j}} \tag{A25}
\end{equation*}
$$

where the value of $\lambda_{j}$, the log-likelihood associated with the $O_{y j}$, depends on whether the $O_{y j}$ are assumed to be normally or lognormally distributed.

When the $O_{y j}$ are normally distributed

$$
\begin{equation*}
\lambda_{j}=-m_{j} \ln q_{j}-\Sigma_{y} \ln E_{y j}-0.5 \Sigma_{y} c_{y j}^{-2}-\Sigma_{y} \ln c_{y j}+\frac{0.5 S_{1}}{q_{j}} \tag{A26}
\end{equation*}
$$

where
(A27) $\quad q_{j}=\frac{\left(S_{1}^{2}+4 m_{j} S_{2}\right)^{0.5}-S_{1}}{2 m_{j}}$
(A28) $S_{1}=\sum_{y}\left(\frac{O_{y j}}{E_{y} c_{y j}}\right)$
(A29) $\quad S_{2}=\sum_{y}\left(\frac{O_{y j}}{E_{y y} c_{y y}}\right)^{2}$
and $m_{j}$ is the number of years for which there is an observation in the $j$ th series of biomass indices (throughout this section all sums over the index $y$ are assumed to cover only these years).

When the $O_{y j}$ are lognormally distributed

$$
\text { (A30) } \lambda_{j}=0.5 m_{j}-\Sigma j \ln \left(O_{t j} s_{x j}\right)-\sum_{y} \frac{0.5}{s_{x j}^{2}}\left[\ln \left(\frac{O_{y j}}{q_{j} E_{y j}}\right)+0.5 s_{y j}^{2}\right]^{2}
$$

where
(A3I) $\quad q_{j}=e^{\left(S_{1}+0.5 m_{j} / / L_{F_{j}}^{-2}\right.}$
(A32) $\quad S_{1}=\sum_{y}\left[s_{x j}-2 \ln \left(\frac{O_{y j}}{E_{y j}}\right)\right]$
and

$$
\begin{equation*}
s_{y j}=\left[\ln \left(1+c_{y y}^{2}\right)\right]^{0.5} \tag{A33}
\end{equation*}
$$

## Fishery Performance Measures

All performance measures are expressed as probabilities (see Tables 28 and 32) and calculated as proportions. Thus, for example, $\mathrm{P}\left(B_{2002}>20 \% B_{0}\right)$ is calculated as the proportion of runs in the importance sample in which the mid-year biomass in 2002 exceeded $20 \%$ of the $B_{0}$ for the run.


[^0]:    * For orange roughy, it is assumed that recruitment to the fishery occurs at maturity, so that $A_{r s}=A_{m s}$ and $S_{r s}=S_{m r}$

